

Package ‘SharpeR’

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Title Statistical Significance of the Sharpe Ratio

BugReports <https://github.com/shabbychef/SharpeR/issues>

Description A collection of tools for analyzing significance of assets, funds, and trading strategies, based on the Sharpe ratio and overfit of the same. Provides density, distribution, quantile and random generation of the Sharpe ratio distribution based on normal returns, as well as the optimal Sharpe ratio over multiple assets. Computes confidence intervals on the Sharpe and provides a test of equality of Sharpe ratios based on the Delta method. The statistical foundations of the Sharpe can be found in the author's Short Sharpe Course <doi:10.2139/ssrn.3036276>.

Depends R (>= 3.0.0)

Imports matrixcalc, methods, sadists (>= 0.2.0)

Suggests xtable, xts, timeSeries, quantmod, MASS, TTR, testthat, sandwich, txtplot, knitr

URL <https://github.com/shabbychef/SharpeR>

VignetteBuilder knitr

Collate 'SharpeR.r' 'data.r' 'utils.r' 'distributions.r' 'sr.r' 'estimation.r' 'sr_bias.r' 'tests.r' 'unified.r'

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 SharpeR-package

statistics concerning Sharpe ratio and Markowitz portfolio

Description

Inference on Sharpe ratio and Markowitz portfolio.

Sharpe Ratio

Suppose x_i are n independent draws of a normal random variable with mean μ and variance σ^2 . Let \bar{x} be the sample mean, and s be the sample standard deviation (using Bessel's correction). Let c_0 be the 'risk free' or 'disastrous rate' of return. Then

$$z = \frac{\bar{x} - c_0}{s}$$

is the (sample) Sharpe ratio.

The units of z are time^{-1/2}. Typically the Sharpe ratio is *annualized* by multiplying by \sqrt{d} , where d is the number of observations per year (or whatever the target annualization epoch.) It is *not* common practice to include units when quoting Sharpe ratio, though doing so could avoid confusion.

The Sharpe ratio follows a rescaled non-central t distribution. That is, z/K follows a non-central t-distribution with m degrees of freedom and non-centrality parameter ζ/K , for some K , m and ζ .

We can generalize Sharpe's model to APT, wherein we write

$$x_i = \alpha + \sum_j \beta_j F_{j,i} + \epsilon_i,$$

where the $F_{j,i}$ are observed 'factor returns', and the variance of the noise term is σ^2 . Via linear regression, one can compute estimates $\hat{\alpha}$, and $\hat{\sigma}$, and then let the 'Sharpe ratio' be

$$z = \frac{\hat{\alpha} - c_0}{\hat{\sigma}}.$$

As above, this Sharpe ratio follows a rescaled t-distribution under normality, *etc.*

The parameters are encoded as follows:

- df stands for the degrees of freedom, typically $n - 1$, but $n - J - 1$ in general.
- ζ is denoted by zeta.
- d is denoted by ope. ('Observations Per Year')
- For the APT form of Sharpe, K stands for the rescaling parameter.

Optimal Sharpe Ratio

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . Let \bar{x} be the (vector) sample mean, and S be the sample covariance matrix (using Bessel's correction). Let

$$Z(w) = \frac{w^\top \bar{x} - c_0}{\sqrt{w^\top S w}}$$

be the (sample) Sharpe ratio of the portfolio w , subject to risk free rate c_0 .

Let w_* be the solution to the portfolio optimization problem:

$$\max_{w: 0 < w^\top S w \leq R^2} Z(w),$$

with maximum value $z_* = Z(w_*)$. Then

$$w_* = R \frac{S^{-1} \bar{x}}{\sqrt{\bar{x}^\top S^{-1} \bar{x}}}$$

and

$$z_* = \sqrt{\bar{x}^\top S^{-1} \bar{x}} - \frac{c_0}{R}$$

The variable z_* follows an *Optimal Sharpe ratio* distribution. For convenience, we may assume that the sample statistic has been annualized in the same manner as the Sharpe ratio, that is by multiplying by d , the number of observations per epoch.

The Optimal Sharpe Ratio distribution is parametrized by the number of assets, q , the number of independent observations, n , the noncentrality parameter,

$$\zeta_* = \sqrt{\mu^\top \Sigma^{-1} \mu},$$

the 'drag' term, c_0/R , and the annualization factor, d . The drag term makes this a location family of distributions, and by default we assume it is zero.

The parameters are encoded as follows:

- q is denoted by df1.
- n is denoted by df2.
- ζ_* is denoted by zeta.s.
- d is denoted by ope.
- c_0/R is denoted by drag.

Spanning and Hedging

As above, let

$$Z(w) = \frac{w^\top \bar{x} - c_0}{\sqrt{w^\top S w}}$$

be the (sample) Sharpe ratio of the portfolio w , subject to risk free rate c_0 .

Let G be a $g \times q$ matrix of 'hedge constraints'. Let w_* be the solution to the portfolio optimization problem:

$$\max_{w: 0 < w^\top S w \leq R^2, G S w = 0} Z(w),$$

with maximum value $z_* = Z(w_*)$. Then z_*^2 can be expressed as the difference of two squared optimal Sharpe ratio random variables. A monotonic transform takes this difference to the LRT statistic for portfolio spanning, first described by Rao, and refined by Giri.

Legal Mumbo Jumbo

SharpeR is distributed in the hope that it will be useful, but WITHOUT ANY WARRANTY; without even the implied warranty of MERCHANTABILITY or FITNESS FOR A PARTICULAR PURPOSE. See the GNU Lesser General Public License for more details.

Note

The following are still in the works:

1. Corrections for standard error based on skew, kurtosis and autocorrelation.
2. Tests on Sharpe under positivity constraint. (*c.f.* Silvapulle)

3. Portfolio spanning tests.
4. Tests on portfolio weights.

This package is maintained as a hobby.

Author(s)

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as.del_sropt

Compute the Sharpe ratio of a hedged Markowitz portfolio.

Description

Computes the Sharpe ratio of the hedged Markowitz portfolio of some observed returns.

Usage

```
as.del_sropt(X, G, drag = 0, ope = 1, epoch = "yr")
```

```
## Default S3 method:
```

```
as.del_sropt(X, G, drag = 0, ope = 1, epoch = "yr")
```

```
## S3 method for class 'xts'
```

```
as.del_sropt(X, G, drag = 0, ope = 1, epoch = "yr")
```

Arguments

X	matrix of returns, or xts object.
G	an $g \times q$ matrix of hedge constraints. A garden variety application would have G be one row of the identity matrix, with a one in the column of the instrument to be 'hedged out'.
drag	the 'drag' term, c_0/R . defaults to 0. It is assumed that drag has been annualized, <i>i.e.</i> has been multiplied by $\sqrt{\text{ope}}$. This is in contrast to the c_0 term given to <code>sr</code> .
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
epoch	the string representation of the 'epoch', defaulting to 'yr'.

Details

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . Let G be a $g \times q$ matrix of rank g . Let \bar{x} be the (vector) sample mean, and S be the sample covariance matrix (using Bessel's correction). Let

$$\zeta(w) = \frac{w^\top \bar{x} - c_0}{\sqrt{w^\top S w}}$$

be the (sample) Sharpe ratio of the portfolio w , subject to risk free rate c_0 .

Let w_* be the solution to the portfolio optimization problem:

$$\max_{w: 0 < w^\top S w \leq R^2, G S w = 0} \zeta(w),$$

with maximum value $z_* = \zeta(w_*)$.

Note that if ope and epoch are not given, the converter from xts attempts to infer the observations per epoch, assuming yearly epoch.

Value

An object of class `del_sropt`.

Author(s)

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See Also

[del_sropt](#), [sropt](#), [sr](#)

Other `del_sropt`: [del_sropt](#), [is.del_sropt\(\)](#)

Examples

```

nfac <- 5
nyr <- 10
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("be deterministic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
# hedge out the first one:
G <- matrix(diag(nfac)[1,],nrow=1)
asro <- as.del_sropt>Returns,G,drag=0,ope=ope)
print(asro)
G <- diag(nfac)[c(1:3),]
asro <- as.del_sropt>Returns,G,drag=0,ope=ope)
# compare to sropt on the remaining assets
# they should be close, but not exact.
asro.alt <- as.sropt>Returns[,4:nfac],drag=0,ope=ope)

# using real data.
if (require(xts)) {
  data(stock_returns)
  # hedge out SPY
  G <- diag(dim(stock_returns)[2])[3,]
  asro <- as.del_sropt(stock_returns,G=G)
}

```

Description

Computes the Sharpe ratio of some observed returns.

Usage

```

as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## Default S3 method:
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## S3 method for class 'matrix'
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## S3 method for class 'data.frame'
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## S3 method for class 'lm'
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## S3 method for class 'xts'
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

## S3 method for class 'timeSeries'
as.sr(x, c0 = 0, ope = 1, na.rm = FALSE, epoch = "yr", higher_order = FALSE)

```

Arguments

x	vector of returns, or object of class data.frame, xts, or lm.
c0	the 'risk-free' or 'disastrous' rate of return. this is assumed to be given in the same units as x, <i>not</i> in 'annualized' terms.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
na.rm	logical. Should missing values be removed?
epoch	the string representation of the 'epoch', defaulting to 'yr'.
higher_order	a Boolean. If true, we compute cumulants of the returns to leverage higher order accuracy formulae when possible.

Details

Suppose x_i are n independent returns of some asset. Let \bar{x} be the sample mean, and s be the sample standard deviation (using Bessel's correction). Let c_0 be the 'risk free rate'. Then

$$z = \frac{\bar{x} - c_0}{s}$$

is the (sample) Sharpe ratio.

The units of z are time^{-1/2}. Typically the Sharpe ratio is *annualized* by multiplying by $\sqrt{\text{ope}}$, where ope is the number of observations per year (or whatever the target annualization epoch.)

Note that if ope is not given, the converter from xts attempts to infer the observations per year, without regard to the name of the epoch given.

Value

a list containing the following components:

sr the annualized Sharpe ratio.
df the t-stat degrees of freedom.
c0 the risk free term.
ope the annualization factor.
rescal the rescaling factor.
epoch the string epoch.
 cast to class sr.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>
 Lo, Andrew W. "The statistics of Sharpe ratios." Financial Analysts Journal 58, no. 4 (2002): 36-52. <http://ssrn.com/paper=377260>

See Also

[reannualize](#)

sr-distribution functions, [dsr](#), [psr](#), [qsr](#), [rsr](#)

Other sr: [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambdap\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
# Sharpe's 'model': just given a bunch of returns.
asr <- as.sr(rnorm(253*3),ope=253)
# or a matrix, with a name
my.returns <- matrix(rnorm(253*3),ncol=1)
colnames(my.returns) <- c("my strategy")
asr <- as.sr(my.returns)

# given an xts object:
if (require(xts)) {
  data(stock_returns)
  IBM <- stock_returns[, 'IBM']
  asr <- as.sr(IBM,na.rm=TRUE)
}

# on a linear model, find the 'Sharpe' of the residual term
```

```

nfac <- 5
nyr <- 10
ope <- 253
set.seed(as.integer(charToRaw("determinstic")))
Factors <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
Betas <- exp(0.1 * rnorm(dim(Factors)[2]))
Returns <- (Factors %*% Betas) + rnorm(dim(Factors)[1],mean=0.0005,sd=0.012)
APT_mod <- lm>Returns ~ Factors)
asr <- as.sr(APT_mod,ope=ope)
# try again, but make the Returns independent of the Factors.
Returns <- rnorm(dim(Factors)[1],mean=0.0005,sd=0.012)
APT_mod <- lm>Returns ~ Factors)
asr <- as.sr(APT_mod,ope=ope)

# compute the Sharpe of a bunch of strategies:
my.returns <- matrix(rnorm(253*3*4),ncol=4)
asr <- as.sr(my.returns) # without sensible colnames?
colnames(my.returns) <- c("strat a","strat b","strat c","strat d")
asr <- as.sr(my.returns)

```

as.sropt

Compute the Sharpe ratio of the Markowitz portfolio.

Description

Computes the Sharpe ratio of the Markowitz portfolio of some observed returns.

Usage

```
as.sropt(X, drag = 0, ope = 1, epoch = "yr")
```

```
## Default S3 method:
```

```
as.sropt(X, drag = 0, ope = 1, epoch = "yr")
```

```
## S3 method for class 'xts'
```

```
as.sropt(X, drag = 0, ope = 1, epoch = "yr")
```

Arguments

X	matrix of returns, or xts object.
drag	the 'drag' term, c_0/R . defaults to 0. It is assumed that drag has been annualized, <i>i.e.</i> has been multiplied by \sqrt{ope} . This is in contrast to the c_0 term given to <code>sr</code> .
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
epoch	the string representation of the 'epoch', defaulting to 'yr'.

Details

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . Let \bar{x} be the (vector) sample mean, and S be the sample covariance matrix (using Bessel's correction). Let

$$\zeta(w) = \frac{w^\top \bar{x} - c_0}{\sqrt{w^\top S w}}$$

be the (sample) Sharpe ratio of the portfolio w , subject to risk free rate c_0 .

Let w_* be the solution to the portfolio optimization problem:

$$\max_{w: 0 < w^\top S w \leq R^2} \zeta(w),$$

with maximum value $z_* = \zeta(w_*)$. Then

$$w_* = R \frac{S^{-1} \bar{x}}{\sqrt{\bar{x}^\top S^{-1} \bar{x}}}$$

and

$$z_* = \sqrt{\bar{x}^\top S^{-1} \bar{x}} - \frac{c_0}{R}$$

The units of z_* are $\text{time}^{-1/2}$. Typically the Sharpe ratio is *annualized* by multiplying by $\sqrt{\text{ope}}$, where ope is the number of observations per year (or whatever the target annualization epoch.)

Note that if ope and epoch are not given, the converter from xts attempts to infer the observations per epoch, assuming yearly epoch.

Value

An object of class `sropt`.

Author(s)

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See Also

[sropt](#), [sr](#), [sropt-distribution functions](#), [dsropt](#), [psropt](#), [qsropt](#), [rsropt](#)

Other `sropt`: [confint.sr\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)

Examples

```
nfac <- 5
nyr <- 10
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("be deterministic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
```

```

# under the alternative:
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0.0005,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
# generating correlated multivariate normal data in a more sane way
if (require(MASS)) {
  nstok <- 10
  nfac <- 3
  nyr <- 10
  ope <- 253
  X.like <- 0.01 * matrix(rnorm(500*nfac),ncol=nfac) %*%
    matrix(runif(nfac*nstok),ncol=nstok)
  Sigma <- cov(X.like) + diag(0.003,nstok)
  # under the null:
  Returns <- mvrnorm(ceiling(ope*nyr),mu=matrix(0,ncol=nstok),Sigma=Sigma)
  asro <- as.sropt>Returns,ope=ope)
  # under the alternative
  Returns <- mvrnorm(ceiling(ope*nyr),mu=matrix(0.001,ncol=nstok),Sigma=Sigma)
  asro <- as.sropt>Returns,ope=ope)
}

# using real data.
if (require(xts)) {
  data(stock_returns)
  asro <- as.sropt(stock_returns)
}

```

confint.sr

Confidence Interval on (optimal) Signal-Noise Ratio

Description

Computes approximate confidence intervals on the (optimal) Signal-Noise ratio given the (optimal) Sharpe ratio. Works on objects of class sr and sropt.

Usage

```

## S3 method for class 'sr'
confint(
  object,
  parm,
  level = 0.95,
  level.lo = (1 - level)/2,
  level.hi = 1 - level.lo,
  type = c("exact", "t", "Z", "Mertens", "Bao"),
  ...
)

## S3 method for class 'sropt'
confint(

```

```

    object,
    parm,
    level = 0.95,
    level.lo = (1 - level)/2,
    level.hi = 1 - level.lo,
    ...
)

## S3 method for class 'del_sropt'
confint(
  object,
  parm,
  level = 0.95,
  level.lo = (1 - level)/2,
  level.hi = 1 - level.lo,
  ...
)

```

Arguments

<code>object</code>	an observed Sharpe ratio statistic, of class <code>sr</code> or <code>sropt</code> .
<code>parm</code>	ignored here, but required for the general method.
<code>level</code>	the confidence level required.
<code>level.lo</code>	the lower confidence level required.
<code>level.hi</code>	the upper confidence level required.
<code>type</code>	which method to apply.
<code>...</code>	further arguments to be passed to or from methods.

Details

Constructs confidence intervals on the Signal-Noise ratio given observed Sharpe ratio statistic. The available methods are:

- exact** The default, which is only exact when returns are normal, based on inverting the non-central t distribution.
- t** Uses the Johnson Welch approximation to the standard error, centered around the sample value.
- Z** Uses the Johnson Welch approximation to the standard error, performing a simple correction for the bias of the Sharpe ratio based on Miller and Gehr formula.
- Mertens** Uses the Mertens higher order approximation to the standard error, centered around the sample value.
- Bao** Uses the Bao higher order approximation to the standard error, performing a higher order correction for the bias of the Sharpe ratio.

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . Let \bar{x} be the (vector) sample mean, and S be the sample covariance matrix (using Bessel's correction). Let

$$z_* = \sqrt{\bar{x}^\top S^{-1} \bar{x}}$$

Given observations of z_* , compute confidence intervals on the population analogue, defined as

$$\zeta_* = \sqrt{\mu^\top \Sigma^{-1} \mu}$$

Value

A matrix (or vector) with columns giving lower and upper confidence limits for the parameter. These will be labelled as level.lo and level.hi in %, e.g. "2.5 %"

Author(s)

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References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

See Also

[confint](#), [se](#)

Other sr: [as.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Other spropt: [as.sropt\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)

Examples

```
# using "sr" class:
ope <- 253
df <- ope * 6
xv <- rnorm(df, 1 / sqrt(ope))
mysr <- as.sr(xv, ope=ope)
confint(mysr, level=0.90)
# using "lm" class
yv <- xv + rnorm(length(xv))
amod <- lm(yv ~ xv)
mysr <- as.sr(amod, ope=ope)
confint(mysr, level.lo=0.05, level.hi=1.0)
# rolling your own.
ope <- 253
df <- ope * 6
zeta <- 1.0
rvs <- rsr(128, df, zeta, ope)
roll.own <- sr(sr=rvs, df=df, c0=0, ope=ope)
aci <- confint(roll.own, level=0.95)
coverage <- 1 - mean((zeta < aci[,1]) | (aci[,2] < zeta))
# using "sropt" class
ope <- 253
```

```

df1 <- 4
df2 <- ope * 3
rvs <- as.matrix(rnorm(df1*df2),ncol=df1)
sro <- as.sropt(rvs,ope=ope)
aci <- confint(sro)
# on sropt, rolling your own.
zeta.s <- 1.0
rvs <- rsropt(128, df1, df2, zeta.s, ope)
roll.own <- sropt(z.s=rvs,df1,df2,drag=0,ope=ope)
aci <- confint(roll.own,level=0.95)
coverage <- 1 - mean((zeta.s < aci[,1]) | (aci[,2] < zeta.s))
# using "del_sropt" class
nfac <- 5
nyr <- 10
ope <- 253
set.seed(as.integer(charToRaw("be deterministic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
# hedge out the first one:
G <- matrix(diag(nfac)[1,],nrow=1)
asro <- as.del_sropt>Returns,G,drag=0,ope=ope)
aci <- confint(asro,level=0.95)
# under the alternative
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0.001,sd=0.0125),ncol=nfac)
asro <- as.del_sropt>Returns,G,drag=0,ope=ope)
aci <- confint(asro,level=0.95)

```

del_sropt

Create an 'del_sropt' object.

Description

Spawns an object of class del_sropt.

Usage

```
del_sropt(z.s, z.sub, df1, df2, df1.sub, drag = 0, ope = 1, epoch = "yr")
```

Arguments

z.s	an optimum Sharpe ratio statistic, on some set of assets.
z.sub	an optimum Sharpe ratio statistic, on a linear subspace of the assets. If larger than z.s an error is thrown.
df1	the number of assets in the portfolio.
df2	the number of observations.
df1.sub	the rank of the linear subspace of the hedge constraint. by restricting attention to the subspace.

drag	the 'drag' term, c_0/R . defaults to 0. It is assumed that drag has been annualized, <i>i.e.</i> has been multiplied by \sqrt{ope} . This is in contrast to the c_0 term given to <code>sr</code> .
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
epoch	the string representation of the 'epoch', defaulting to 'yr'.

Details

The `del_sropt` class contains information about the difference between two rescaled T^2 -statistics, useful for spanning tests, and inference on hedged portfolios. The following are list attributes of the object:

sropt The (optimal) Sharpe ratio statistic of the 'full' set of assets.

sropt.sub The (optimal) Sharpe ratio statistic on some subset, or linear subspace, of the assets.

df1 The number of assets.

df2 The number of observations.

df1.sub The number of degrees of freedom in the hedge constraint.

drag The drag term, which is the 'risk free rate' divided by the maximum risk.

ope The 'observations per epoch'.

epoch The string name of the 'epoch'.

For the most part, this constructor should *not* be called directly, rather `as.del_sropt` should be called instead to compute the needed statistics.

Value

a list cast to class `del_sropt`, with attributes

sropt the optimal Sharpe statistic.

sropt.sub the optimal Sharpe statistic on the subspace.

df1 the number of assets.

df2 the number of observed vectors.

df1.sub the input `df1.sub` term.

drag the input drag term.

ope the input ope term.

T2 the Hotelling T^2 statistic.

T2.sub the Hotelling T^2 statistic on the subspace.

Note

WARNING: This function is not well tested, may contain errors, may change in the next package update. Take caution.

2FIX: allow rownames?

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

[reannualize](#)

[as.del_sropt](#)

Other del_sropt: [as.del_sropt\(\)](#), [is.del_sropt\(\)](#)

Examples

```
# roll your own.
ope <- 253

set.seed(as.integer(charToRaw("be deterministic")))
n.stock <- 10
X <- matrix(rnorm(1000*n.stock),nrow=1000)
Sigma <- cov(X)
mu <- colMeans(X)
w <- solve(Sigma,mu)
z <- t(mu) %*% w
n.sub <- 6
w.sub <- solve(Sigma[1:n.sub,1:n.sub],mu[1:n.sub])
z.sub <- t(mu[1:n.sub]) %*% w.sub
df1.sub <- n.stock - n.sub

roll.own <- del_sropt(z.s=z,z.sub=z.sub,df1=10,df2=1000,
  df1.sub=df1.sub,ope=ope)
print(roll.own)
```

 dsr

The (non-central) Sharpe ratio.

Description

Density, distribution function, quantile function and random generation for the Sharpe ratio distribution with df degrees of freedom (and optional signal-noise-ratio zeta).

Usage

```
dsr(x, df, zeta, ope, ...)
```

```
psr(q, df, zeta, ope, ...)
```

```
qsr(p, df, zeta, ope, ...)
```

```
rsr(n, df, zeta, ope)
```

Arguments

x, q	vector of quantiles.
df	the number of observations the statistic is based on. This is one more than the number of degrees of freedom in the corresponding t-statistic, although the effect will be small when df is large.
zeta	the 'signal-to-noise' parameter, ζ defined as the population mean divided by the population standard deviation, 'annualized'.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
...	arguments passed on to the respective t-distribution functions, namely <code>lower.tail</code> with default TRUE, <code>log</code> with default FALSE, and <code>log.p</code> with default FALSE.
p	vector of probabilities.
n	number of observations.

Details

Suppose x_i are n independent draws of a normal random variable with mean μ and variance σ^2 . Let \bar{x} be the sample mean, and s be the sample standard deviation (using Bessel's correction). Let c_0 be the 'risk free rate'. Then

$$z = \frac{\bar{x} - c_0}{s}$$

is the (sample) Sharpe ratio.

The units of z is time^{-1/2}. Typically the Sharpe ratio is *annualized* by multiplying by \sqrt{d} , where d is the number of observations per epoch (typically a year).

Letting $z = \sqrt{d} \frac{\bar{x} - c_0}{s}$, where the sample estimates are based on n observations, then z takes a (non-central) Sharpe ratio distribution parametrized by n 'degrees of freedom', non-centrality parameter $\zeta = \frac{\mu - c_0}{\sigma}$, and annualization parameter d .

The parameters are encoded as follows:

- n is denoted by df.
- ζ is denoted by zeta.
- d is denoted by ope. ('Observations Per Year')

If the returns violate the assumptions of normality, independence, etc (*as they always should in the real world*), the sample Sharpe Ratio will not follow this distribution. It does provide, however, a reasonable approximation in many cases.

Value

dsr gives the density, psr gives the distribution function, qsr gives the quantile function, and rsr generates random deviates.

Invalid arguments will result in return value NaN with a warning.

Note

This is a thin wrapper on the t distribution. The functions `dt`, `pt`, `qt` can accept `ncp` from limited range ($|\delta| \leq 37.62$). Some corrections may have to be made here for large zeta.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

See Also

[reannualize](#)

t-distribution functions, [dt](#), [pt](#), [qt](#), [rt](#)

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
rvs <- rsr(128, 253*6, 0, 253)
dvs <- dsr(rvs, 253*6, 0, 253)
pvs.H0 <- psr(rvs, 253*6, 0, 253)
pvs.HA <- psr(rvs, 253*6, 1, 253)

plot(ecdf(pvs.H0))
plot(ecdf(pvs.HA))
```

dsropt

The (non-central) maximal Sharpe ratio distribution.

Description

Density, distribution function, quantile function and random generation for the maximal Sharpe ratio distribution with `df1` and `df2` degrees of freedom (and optional maximal signal-noise-ratio `zeta.s`).

Usage

dsropt(x, df1, df2, zeta.s, ope, drag = 0, log = FALSE)

psropt(q, df1, df2, zeta.s, ope, drag = 0, ...)

qsropt(p, df1, df2, zeta.s, ope, drag = 0, ...)

rsropt(n, df1, df2, zeta.s, ope, drag = 0, ...)

Arguments

x, q	vector of quantiles.
df1	the number of assets in the portfolio.
df2	the number of observations.
zeta.s	the non-centrality parameter, defined as $\zeta_* = \sqrt{\mu^\top \Sigma^{-1} \mu}$, for population parameters. defaults to 0, <i>i.e.</i> a central maximal Sharpe ratio distribution.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
drag	the 'drag' term, c_0/R . defaults to 0. It is assumed that drag has been annualized, <i>i.e.</i> is given in the same units as x and q.
log	logical; if TRUE, densities f are given as $\log(f)$.
p	vector of probabilities.
n	number of observations.
...	arguments passed on to the respective Hotelling T^2 functions.

Details

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . Let \bar{x} be the (vector) sample mean, and S be the sample covariance matrix (using Bessel's correction). Let

$$Z(w) = \frac{w^\top \bar{x} - c_0}{\sqrt{w^\top S w}}$$

be the (sample) Sharpe ratio of the portfolio w , subject to risk free rate c_0 .

Let w_* be the solution to the portfolio optimization problem:

$$\max_{w: 0 < w^\top S w \leq R^2} Z(w),$$

with maximum value $z_* = Z(w_*)$. Then

$$w_* = R \frac{S^{-1} \bar{x}}{\sqrt{\bar{x}^\top S^{-1} \bar{x}}}$$

and

$$z_* = \sqrt{\bar{x}^\top S^{-1} \bar{x}} - \frac{c_0}{R}$$

The variable z_* follows an *Optimal Sharpe ratio* distribution. For convenience, we may assume that the sample statistic has been annualized in the same manner as the Sharpe ratio, that is by multiplying by d , the number of observations per epoch.

The Optimal Sharpe Ratio distribution is parametrized by the number of assets, q , the number of independent observations, n , the noncentrality parameter,

$$\zeta_* = \sqrt{\mu^\top \Sigma^{-1} \mu},$$

the 'drag' term, c_0/R , and the annualization factor, d . The drag term makes this a location family of distributions, and by default we assume it is zero.

The parameters are encoded as follows:

- q is denoted by `df1`.
- n is denoted by `df2`.
- ζ_* is denoted by `zeta.s`.
- d is denoted by `ope`.
- c_0/R is denoted by `drag`.

Value

`dsropt` gives the density, `psropt` gives the distribution function, `qsropt` gives the quantile function, and `rsropt` generates random deviates.

Invalid arguments will result in return value NaN with a warning.

Note

This is a thin wrapper on the Hotelling T-squared distribution, which is a wrapper on the F distribution.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Kan, Raymond and Smith, Daniel R. "The Distribution of the Sample Minimum-Variance Frontier." *Journal of Management Science* 54, no. 7 (2008): 1364–1380. <http://mansci.journal.informs.org/cgi/doi/10.1287/mnsc.1070.0852>

See Also

[reannualize](#)

F-distribution functions, [df](#), [pf](#), [qf](#), [rf](#), Sharpe ratio distribution, [dsr](#), [psr](#), [qsr](#), [rsr](#).

Other spropt: [as.sropt\(\)](#), [confint.sr\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)

Examples

```
# generate some variates
ngen <- 128
ope <- 253
df1 <- 8
df2 <- ope * 10
drag <- 0
# sample
rvs <- rsropt(ngen, df1, df2, drag, ope)
hist(rvs)
# these should be uniform:
isp <- psropt(rvs, df1, df2, drag, ope)
plot(ecdf(isp))
```

inference

Inference on noncentrality parameter of F-like statistic

Description

Estimates the non-centrality parameter associated with an observed statistic following an optimal Sharpe Ratio distribution.

Usage

```
inference(z.s, type = c("KRS", "MLE", "unbiased"))

## S3 method for class 'sropt'
inference(z.s, type = c("KRS", "MLE", "unbiased"))

## S3 method for class 'del_sropt'
inference(z.s, type = c("KRS", "MLE", "unbiased"))
```

Arguments

`z.s` an object of type `sropt`, or `del_sropt`
`type` the estimator type. one of `c("KRS", "MLE", "unbiased")`

Details

Let F be an observed statistic distributed as a non-central F with ν_1, ν_2 degrees of freedom and non-centrality parameter δ^2 . Three methods are presented to estimate the non-centrality parameter from the statistic:

- an unbiased estimator, which, unfortunately, may be negative.
- the Maximum Likelihood Estimator, which may be zero, but not negative.
- the estimator of Kubokawa, Roberts, and Shaleh (KRS), which is a shrinkage estimator.

The spropt distribution is equivalent to an F distribution up to a square root and some rescalings.

The non-centrality parameter of the spropt distribution is the square root of that of the Hotelling, *i.e.* has units 'per square root time'. As such, the 'unbiased' type can be problematic!

Value

an estimate of the non-centrality parameter, which is the maximal population Sharpe ratio.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Kubokawa, T., C. P. Robert, and A. K. Saleh. "Estimation of noncentrality parameters." *Canadian Journal of Statistics* 21, no. 1 (1993): 45-57. <https://www.jstor.org/stable/3315657>

Spruill, M. C. "Computation of the maximum likelihood estimate of a noncentrality parameter." *Journal of multivariate analysis* 18, no. 2 (1986): 216-224. <http://www.sciencedirect.com/science/article/pii/0047259X86900709>

See Also

F-distribution functions, [df](#).

Other spropt Hotelling: [sric\(\)](#)

Examples

```
# generate some spropts
nfac <- 3
nyr <- 5
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("determinstic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
est1 <- inference(asro,type='unbiased')
est2 <- inference(asro,type='KRS')
est3 <- inference(asro,type='MLE')

# under the alternative:
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0.0005,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
est1 <- inference(asro,type='unbiased')
est2 <- inference(asro,type='KRS')
est3 <- inference(asro,type='MLE')

# sample many under the alternative, look at the estimator.
df1 <- 3
df2 <- 512
ope <- 253
```

```
zeta.s <- 1.25
rvs <- rsropt(128, df1, df2, zeta.s, ope)
roll.own <- sropt(z.s=rvs,df1,df2,drag=0,ope=ope)
est1 <- inference(roll.own,type='unbiased')
est2 <- inference(roll.own,type='KRS')
est3 <- inference(roll.own,type='MLE')

# for del_sropt:
nfac <- 5
nyr <- 10
ope <- 253
set.seed(as.integer(charToRaw("fix seed")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0.0005,sd=0.0125),ncol=nfac)
# hedge out the first one:
G <- matrix(diag(nfac)[1,],nrow=1)
asro <- as.del_sropt>Returns,G,drag=0,ope=ope)
est1 <- inference(asro,type='unbiased')
est2 <- inference(asro,type='KRS')
est3 <- inference(asro,type='MLE')
```

is.del_sropt

Is this in the "del_sropt" class?

Description

Checks if an object is in the class 'del_sropt'

Usage

```
is.del_sropt(x)
```

Arguments

x an object of some kind.

Details

To satisfy the minimum requirements of an S3 class.

Value

a boolean.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

del_sropt

Other del_sropt: [as.del_sropt\(\)](#), [del_sropt](#)

Examples

```
roll.own <- del_sropt(z.s=2,z.sub=1,df1=10,df2=1000,df1.sub=3,ope=1,epoch="yr")
is.sropt(roll.own)
```

is.sr

Is this in the "sr" class?

Description

Checks if an object is in the class 'sr'

Usage

```
is.sr(x)
```

Arguments

x an object of some kind.

Details

To satisfy the minimum requirements of an S3 class.

Value

a boolean.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

sr

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [plambdap\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr](#), [summary.sr](#)

Examples

```
rvs <- as.sr(rnorm(253*8),ope=253)
is.sr(rvs)
```

is.sropt

Is this in the "sropt" class?

Description

Checks if an object is in the class 'sropt'

Usage

```
is.sropt(x)
```

Arguments

x an object of some kind.

Details

To satisfy the minimum requirements of an S3 class.

Value

a boolean.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

sropt

Other sropt: [as.sropt\(\)](#), [confint.sr\(\)](#), [dsropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)

Examples

```
nfac <- 5
nyr <- 10
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("be deterministic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
is.sropt(asro)
```

ism_vcov *Compute variance covariance of Inverse 'Unified' Second Moment*

Description

Computes the variance covariance matrix of the inverse unified second moment matrix.

Usage

```
ism_vcov(X,vcov.func=vcov,fit.intercept=TRUE)
```

Arguments

`X` an $n \times p$ matrix of observed returns.

`vcov.func` a function which takes an object of class `lm`, and computes a variance-covariance matrix. If equal to the string "normal", we assume multivariate normal returns.

`fit.intercept` a boolean controlling whether we add a column of ones to the data, or fit the raw uncentered second moment.

Details

Given p -vector x with mean μ and covariance, Σ , let y be x with a one prepended. Then let $\Theta = E(yy^\top)$, the uncentered second moment matrix. The inverse of Θ contains the (negative) Markowitz portfolio and the precision matrix.

Given n contemporaneous observations of p -vectors, stacked as rows in the $n \times p$ matrix X , this function estimates the mean and the asymptotic variance-covariance matrix of Θ^{-1} .

One may use the default method for computing covariance, via the `vcov` function, or via a 'fancy' estimator, like `sandwich:vcovHAC`, `sandwich:vcovHC`, *etc.*

Value

a list containing the following components:

`mu` a $q = p(p + 3)/2$ vector of the negative Markowitz portfolio, then the vech'd precision matrix of the sample data

`Ohat` the $q \times q$ estimated variance covariance matrix.

`n` the number of rows in X .

`p` the number of assets.

Note

By flipping the sign of X , the inverse of Θ contains the *positive* Markowitz portfolio and the precision matrix on X . Performing this transform before passing the data to this function should be considered idiomatic.

This function will be deprecated in future releases of this package. Users should migrate at that time to a similar function in the MarkowitzR package.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Pav, S. E. "Asymptotic Distribution of the Markowitz Portfolio." 2013 <http://arxiv.org/abs/1312.0557>

See Also

[sm_vcov](#), [sr_vcov](#)

Examples

```
X <- matrix(rnorm(1000*3),ncol=3)
# putting in -X is idiomatic:
ism <- ism_vcov(-X)
iSigmas.n <- ism_vcov(-X,vcov.func="normal")
iSigmas.n <- ism_vcov(-X,fit.intercept=FALSE)
# compute the marginal Wald test statistics:
ism.mu <- ism$mu[1:ism$p]
ism.Sg <- ism$Ohat[1:ism$p,1:ism$p]
wald.stats <- ism.mu / sqrt(diag(ism.Sg))

# make it fat tailed:
X <- matrix(rt(1000*3,df=5),ncol=3)
ism <- ism_vcov(X)
wald.stats <- ism$mu[1:ism$p] / sqrt(diag(ism$Ohat[1:ism$p,1:ism$p]))

if (require(sandwich)) {
  ism <- ism_vcov(X,vcov.func=vcovHC)
  wald.stats <- ism$mu[1:ism$p] / sqrt(diag(ism$Ohat[1:ism$p,1:ism$p]))
}

# add some autocorrelation to X
Xf <- filter(X,c(0.2),"recursive")
colnames(Xf) <- colnames(X)
ism <- ism_vcov(Xf)
wald.stats <- ism$mu[1:ism$p] / sqrt(diag(ism$Ohat[1:ism$p,1:ism$p]))

if (require(sandwich)) {
  ism <- ism_vcov(Xf,vcov.func=vcovHAC)
  wald.stats <- ism$mu[1:ism$p] / sqrt(diag(ism$Ohat[1:ism$p,1:ism$p]))
}
```

pco_sropt

*The 'confidence distribution' for maximal Sharpe ratio.***Description**

Distribution function and quantile function for the 'confidence distribution' of the maximal Sharpe ratio. This is just an inversion to perform inference on ζ_* given observed statistic z_* .

Usage

```
pco_sropt(q, df1, df2, z.s, ope, lower.tail=TRUE, log.p=FALSE)
```

```
qco_sropt(p, df1, df2, z.s, ope, lower.tail=TRUE, log.p=FALSE, lb=0, ub=Inf)
```

Arguments

q	vector of quantiles.
df1	the number of assets in the portfolio.
df2	the number of observations.
z.s	an observed Sharpe ratio statistic, annualized.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$, otherwise, $P[X > x]$.
log.p	logical; if TRUE, probabilities p are given as $\log(p)$.
p	vector of probabilities.
lb	the lower bound for the output of qco_sropt.
ub	the upper bound for the output of qco_sropt.

Details

Suppose z_* follows a *Maximal Sharpe ratio* distribution (see [SharpeR-package](#)) for known degrees of freedom, and unknown non-centrality parameter ζ_* . The 'confidence distribution' views ζ_* as a random quantity once z_* is observed. As such, the CDF of the confidence distribution is the same as that of the Maximal Sharpe ratio (up to a flip of `lower.tail`); while the quantile function is used to compute confidence intervals on ζ_* given z_* .

Value

pco_sropt gives the distribution function, and qco_sropt gives the quantile function.

Invalid arguments will result in return value NaN with a warning.

Note

When `lower.tail` is true, `pco_sropt` is monotonic increasing with respect to `q`, and decreasing in `sropt`; these are reversed when `lower.tail` is false. Similarly, `qco_sropt` is increasing in $\text{sign}(\text{as.double}(\text{lower.tail}) - 0.5) * p$ and $-\text{sign}(\text{as.double}(\text{lower.tail}) - 0.5) * \text{sropt}$.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

[reannualize](#)

[dsropt](#), [psropt](#), [qsropt](#), [rsropt](#)

Other `sropt`: [as.sropt\(\)](#), [confint.sr\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)

Examples

```
zeta.s <- 2.0
ope <- 253
ntest <- 50
df1 <- 4
df2 <- 6 * ope
rvs <- rsropt(ntest, df1=df1, df2=df2, zeta.s=zeta.s)
qvs <- seq(0, 10, length.out=51)
pps <- pco_sropt(qvs, df1, df2, rvs[1], ope)

if (require(txtplot))
  txtplot(qvs, pps)

pps <- pco_sropt(qvs, df1, df2, rvs[1], ope, lower.tail=FALSE)

if (require(txtplot))
  txtplot(qvs, pps)

svs <- seq(0, 4, length.out=51)
pps <- pco_sropt(2, df1, df2, svs, ope)
pps <- pco_sropt(2, df1, df2, svs, ope, lower.tail=FALSE)

pps <- pco_sropt(qvs, df1, df2, rvs[1], ope, lower.tail=FALSE)
pco_sropt(-1, df1, df2, rvs[1], ope)

qvs <- qco_sropt(0.05, df1=df1, df2=df2, z.s=rvs)
mean(qvs > zeta.s)
qvs <- qco_sropt(0.5, df1=df1, df2=df2, z.s=rvs)
mean(qvs > zeta.s)
qvs <- qco_sropt(0.95, df1=df1, df2=df2, z.s=rvs)
mean(qvs > zeta.s)
# test vectorization:
```

```

qv <- qco_sropt(0.1,df1,df2,rvs)
qv <- qco_sropt(c(0.1,0.2),df1,df2,rvs)
qv <- qco_sropt(c(0.1,0.2),c(df1,2*df1),df2,rvs)
qv <- qco_sropt(c(0.1,0.2),c(df1,2*df1),c(df2,2*df2),rvs)

```

plambdap

The lambda-prime distribution.

Description

Distribution function and quantile function for LeCoutre's lambda-prime distribution with df degrees of freedom and the observed t-statistic, $tstat$.

Usage

```
plambdap(q, df, tstat, lower.tail = TRUE, log.p = FALSE)
```

```
qlambdap(p, df, tstat, lower.tail = TRUE, log.p = FALSE)
```

```
rlambdap(n, df, tstat)
```

Arguments

<code>q</code>	vector of quantiles.
<code>df</code>	the degrees of freedom of the t-statistic.
<code>tstat</code>	the observed (non-central) t-statistic.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$, otherwise, $P[X > x]$.
<code>log.p</code>	logical; if TRUE, probabilities p are given as $\log(p)$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If 'length(n) > 1', the length is taken to be the number required.

Details

Let t be distributed as a non-central t with ν degrees of freedom and non-centrality parameter δ . We can view this as

$$t = \frac{Z + \delta}{\sqrt{V/\nu}}$$

where Z is a standard normal, δ is the non-centrality parameter, V is a chi-square RV with ν degrees of freedom, independent of Z . We can rewrite this as

$$\delta = t\sqrt{V/\nu} + Z.$$

Thus a 'lambda-prime' random variable with parameters t and ν is one expressible as a sum

$$t\sqrt{V/\nu} + Z$$

for Chi-square V with ν d.f., independent from standard normal Z

Value

d`lambdap` gives the density, p`lambdap` gives the distribution function, q`lambdap` gives the quantile function, and r`lambdap` generates random deviates.

Invalid arguments will result in return value NaN with a warning.

Note

p`lambdap` should be an increasing function of the argument `q`, and decreasing in `tstat`. q`lambdap` should be increasing in `p`

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Lecoutre, Bruno. "Another look at confidence intervals for the noncentral t distribution." *Journal of Modern Applied Statistical Methods* 6, no. 1 (2007): 107–116. [http://eris62.eu/telechargements/Lecoutre_Another_look-JMSAM2007_6\(1\).pdf](http://eris62.eu/telechargements/Lecoutre_Another_look-JMSAM2007_6(1).pdf)

Lecoutre, Bruno. "Two useful distributions for Bayesian predictive procedures under normal models." *Journal of Statistical Planning and Inference* 79 (1999): 93–105.

See Also

t-distribution functions, [dt](#), [pt](#), [qt](#), [rt](#)

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
rvs <- rnorm(128)
pvs <- plambdap(rvs, 253*6, 0.5)
plot(ecdf(pvs))
pvs <- plambdap(rvs, 253*6, 1)
plot(ecdf(pvs))
pvs <- plambdap(rvs, 253*6, -0.5)
plot(ecdf(pvs))
# test vectorization:
qv <- qlambdap(0.1,128,2)
qv <- qlambdap(c(0.1),128,2)
qv <- qlambdap(c(0.2),128,2)
qv <- qlambdap(c(0.2),253,2)
qv <- qlambdap(c(0.1,0.2),128,2)
qv <- qlambdap(c(0.1,0.2),c(128,253),2)
qv <- qlambdap(c(0.1,0.2),c(128,253),c(2,4))
qv <- qlambdap(c(0.1,0.2),c(128,253),c(2,4,8,16))
# random generation
rv <- rlambdap(1000,252,2)
```

power.sropt_test *Power calculations for optimal Sharpe ratio tests*

Description

Compute power of test, or determine parameters to obtain target power.

Usage

```
power.sropt_test(df1=NULL,df2=NULL,zeta.s=NULL,
                 sig.level=0.05,power=NULL,ope=1)
```

Arguments

df1	the number of assets in the portfolio.
df2	the number of observations.
zeta.s	the 'signal-to-noise' parameter, defined as ...
sig.level	Significance level (Type I error probability).
power	Power of test (1 minus Type II error probability).
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.

Details

Suppose you perform a single-sample test for significance of the optimal Sharpe ratio based on the corresponding single-sample T^2 -test. Given any four of: the effect size (the population optimal SNR, ζ_*), the number of assets, the number of observations, and the type I and type II rates, this function computes the fifth.

Exactly one of the parameters df1, df2, zeta.s, power, and sig.level must be passed as NULL, and that parameter is determined from the others. Notice that sig.level has non-NULL default, so NULL must be explicitly passed if you want to compute it.

Value

Object of class power.htest, a list of the arguments (including the computed one) augmented with method, note and n.epoch elements, the latter is the number of epochs under the given annualization (ope), NA if none given.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also[reannualize](#)[power.t.test](#), [sropt_test](#)Other sropt: [as.sropt\(\)](#), [confint.sr\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#), [sropt](#)**Examples**

```
anex <- power.sropt_test(8, 4*253, 1, 0.05, NULL, ope=253)
```

power.sr_test

Power calculations for Sharpe ratio tests

Description

Compute power of test, or determine parameters to obtain target power.

Usage

```
power.sr_test(n=NULL, zeta=NULL, sig.level=0.05, power=NULL,
              alternative=c("one.sided", "two.sided"), ope=NULL)
```

Arguments

n	Number of observations
zeta	the 'signal-to-noise' parameter, defined as the population mean divided by the population standard deviation, 'annualized'.
sig.level	Significance level (Type I error probability).
power	Power of test (1 minus Type II error probability).
alternative	One- or two-sided test.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.

Details

Suppose you perform a single-sample test for significance of the Sharpe ratio based on the corresponding single-sample t-test. Given any three of: the effect size (the population SNR, ζ), the number of observations, and the type I and type II rates, this function computes the fourth.

This is a thin wrapper on [power.t.test](#).

Exactly one of the parameters n, zeta, power, and sig.level must be passed as NULL, and that parameter is determined from the others. Notice that sig.level has non-NULL default, so NULL must be explicitly passed if you want to compute it.

Value

Object of class `power.htest`, a list of the arguments (including the computed one) augmented with `method`, `note` and `n.epoch` elements, the latter is the number of epochs under the given annualization (`ope`), NA if none given.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." *Journal of business* (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

Johnson, N. L., and Welch, B. L. "Applications of the non-central t-distribution." *Biometrika* 31, no. 3-4 (1940): 362-389. <http://dx.doi.org/10.1093/biomet/31.3-4.362>

Lehr, R. "Sixteen S-squared over D-squared: A relation for crude sample size estimates." *Statist. Med.*, 11, no 8 (1992): 1099–1102. doi: 10.1002/sim.4780110811

See Also

[reannualize](#)

[power.t.test](#), [sr_test](#)

Other `sr`: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
anex <- power.sr_test(253,1,0.05,NULL,ope=253)
anex <- power.sr_test(n=253,zeta=NULL,sig.level=0.05,power=0.5,ope=253)
anex <- power.sr_test(n=NULL,zeta=0.6,sig.level=0.05,power=0.5,ope=253)
# Lehr's Rule
zetas <- seq(0.1,2.5,length.out=51)
ssizes <- sapply(zetas,function(zed) {
  x <- power.sr_test(n=NULL,zeta=zed,sig.level=0.05,power=0.8,
    alternative="two.sided",ope=253)
  x$n / 253})
# should be around 8.
print(summary(ssizes * zetas * zetas))
# e = n z^2 mnemonic approximate rule for 0.05 type I, 50% power
ssizes <- sapply(zetas,function(zed) {
  x <- power.sr_test(n=NULL,zeta=zed,sig.level=0.05,power=0.5,ope=253)
  x$n / 253 })
print(summary(ssizes * zetas * zetas - exp(1)))
```

predint	<i>prediction interval for Sharpe ratio</i>
---------	---

Description

Computes the prediction interval for Sharpe ratio.

Usage

```
predint(x, oosdf, oosrescal=1/sqrt(oosdf+1), ope=NULL, level=0.95,
        level.lo=(1-level)/2, level.hi=1-level.lo)
```

Arguments

x	a (non-empty) numeric vector of data values, or an object of class <code>sr</code> .
oosdf	the future (or 'out of sample', thus 'oos') degrees of freedom. In the vanilla Sharpe case, this is the number of future observations <i>minus one</i> .
oosrescal	the rescaling parameter for the future Sharpe ratio. The default value holds for the case of unattributed models ('vanilla Sharpe'), but can be set to some other value to deal with the magnitude of attribution factors in the future period.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of <code>ope</code> per epoch. The default value is to take the same <code>ope</code> from the input <code>x</code> object, if it is unambiguous.
level	the confidence level required.
level.lo	the lower confidence level required.
level.hi	the upper confidence level required.

Details

Given n_0 observations x_i from a normal random variable, with mean μ and standard deviation σ , computes an interval $[y_1, y_2]$ such that with a fixed probability, the sample Sharpe ratio over n future observations will fall in the given interval. The coverage is over repeated draws of both the past and future data, thus this computation takes into account error in both the estimate of Sharpe and the as yet unrealized returns.

Value

A matrix (or vector) with columns giving lower and upper confidence limits for the parameter. These will be labelled as `level.lo` and `level.hi` in %, e.g. "2.5 %"

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

Pav, Steven. "Inference on the Sharpe ratio via the upsilon distribution." Arxiv (2015). <http://arxiv.org/abs/1505.00829>

See Also

[confint.sr](#).

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
# should reject null
etc <- predint(rnorm(1000,mean=0.5,sd=0.1),oosdf=127,ope=1)
etc <- predint(matrix(rnorm(1000*5,mean=0.05),ncol=5),oosdf=63,ope=1)

# check coverage
mu <- 0.0005
sg <- 0.013
n1 <- 512
n2 <- 256
p <- 100
x1 <- matrix(rnorm(n1*p,mean=mu,sd=sg),ncol=p)
x2 <- matrix(rnorm(n2*p,mean=mu,sd=sg),ncol=p)
sr1 <- as.sr(x1)
sr2 <- as.sr(x2)
# check coverage of prediction interval
etc1 <- predint(sr1,oosdf=n2-1,level=0.95)
is.ok <- (etc1[,1] <= sr2$sr) & (sr2$sr <= etc1[,2])
covr <- mean(is.ok)
```

print.sr

Print values.

Description

Displays an object, returning it *invisibly*, (via `invisible(x)`.)

Usage

```
## S3 method for class 'sr'
print(x, ...)

## S3 method for class 'sropt'
print(x, ...)

## S3 method for class 'del_sropt'
print(x, ...)
```

Arguments

x an object of class sr or sropt.
... further arguments to be passed to or from methods.

Value

the object, wrapped in invisible.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

See Also

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambdap\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr_summary.sr](#)

Examples

```
# compute a 'daily' Sharpe
mysr <- as.sr(rnorm(253*8), ope=1, epoch="day")
print(mysr)
# roll your own.
ope <- 253
zeta <- 1.0
n <- 6 * ope
rvs <- rsr(1, n, zeta, ope=ope)
roll.own <- sr(sr=rvs, df=n-1, ope=ope, rescal=sqrt(1/n))
print(roll.own)
# put a bunch in. naming becomes a problem.
rvs <- rsr(5, n, zeta, ope=ope)
roll.own <- sr(sr=rvs, df=n-1, ope=ope, rescal=sqrt(1/n))
print(roll.own)
```

```
# for spropt objects:
nfac <- 5
nyr <- 10
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("be deterministic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
print(asro)
```

reannualize

Change the annualization of a Sharpe ratio.

Description

Changes the annualization factor of a Sharpe ratio statistic, or the rate at which observations are made.

Usage

```
reannualize(object, new.ope = NULL, new.epoch = NULL)
```

```
## S3 method for class 'sr'
reannualize(object, new.ope = NULL, new.epoch = NULL)
```

```
## S3 method for class 'spropt'
reannualize(object, new.ope = NULL, new.epoch = NULL)
```

Arguments

object	an object of class sr or spropt.
new.ope	the new observations per epoch. If none given, it is not updated.
new.epoch	a string representation of the epoch. If none given, it is not updated.

Value

the input object with the annualization and/or epoch updated.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

sr

sropt

Other sr: `as.sr()`, `confint.sr()`, `dssr()`, `is.sr()`, `plambda()`, `power.sr_test()`, `predint()`, `print.sr()`, `se()`, `sr_equality_test()`, `sr_test()`, `sr_unpaired_test()`, `sr_vcov()`, `sr_summary.sr`

Other sropt: `as.sropt()`, `confint.sr()`, `dsropt()`, `is.sropt()`, `pco_sropt()`, `power.sropt_test()`, `sropt_test()`, `sropt`

Examples

```
# compute a 'daily' Sharpe
mysr <- as.sr(rnorm(253*8), ope=1, epoch="day")
# turn into annual
mysr2 <- reannualize(mysr, new.ope=253, new.epoch="yr")

# for sropt
ope <- 253
zeta.s <- 1.0
df1 <- 10
df2 <- 6 * ope
rvs <- rsropt(1, df1, df2, zeta.s, ope, drag=0)
roll.own <- sropt(z.s=rvs, df1, df2, drag=0, ope=ope, epoch="yr")
# make 'monthly'
roll.monthly <- reannualize(roll.own, new.ope=21, new.epoch="mo.")
# make 'daily'
roll.daily <- reannualize(roll.own, new.ope=1, new.epoch="day")
```

se

Standard error computation

Description

Estimates the standard error of the Sharpe ratio statistic.

Usage

```
se(z, type)

## S3 method for class 'sr'
se(z, type = c("t", "Lo", "Mertens", "Bao"))
```

Arguments

z an observed Sharpe ratio statistic, of class sr.
type estimator type. one of "t", "Lo", "Mertens", "Bao"

Details

For an observed Sharpe ratio, estimate the standard error. The following methods are recognized:

t The default, based on Johnson & Welch, with a correction for small sample size. Also known as 'Lo'.

Mertens An approximation to the standard error taking into skewness and kurtosis of the returns distribution.

Bao An even higher accuracy approximation using higher order moments.

There should be very little difference between these except for very small sample sizes.

Value

an estimate of standard error.

Note

The units of the standard error are consistent with those of the input `sr` object.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." *Journal of business* (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

Johnson, N. L., and Welch, B. L. "Applications of the non-central t-distribution." *Biometrika* 31, no. 3-4 (1940): 362-389. <http://dx.doi.org/10.1093/biomet/31.3-4.362>

Lo, Andrew W. "The statistics of Sharpe ratios." *Financial Analysts Journal* 58, no. 4 (2002): 36-52. <http://ssrn.com/paper=377260>

Bao, Yong. "Estimation Risk-Adjusted Sharpe Ratio and Fund Performance Ranking Under a General Return Distribution." *Journal of Financial Econometrics* 7, no. 2 (2009): 152-173. <https://doi.org/10.1093/jfinec/nbn022>

Opdyke, J. D. "Comparing Sharpe Ratios: So Where are the p-values?" *Journal of Asset Management* 8, no. 5 (2006): 308-336. <http://ssrn.com/paper=886728>

Walck, C. "Hand-book on STATISTICAL DISTRIBUTIONS for experimentalists." 1996. <http://www.stat.rice.edu/~dobelman/textfiles/DistributionsHandbook.pdf>

See Also

sr-distribution functions, [dsr](#), [sr_variance](#).

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambdap\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr](#), [summary.sr](#)

Examples

```
asr <- as.sr(rnorm(128,0.2))
anse <- se(asr,type="t")
anse <- se(asr,type="Lo")
```

SharpeR-NEWS

News for package 'SharpeR':

Description

News for package 'SharpeR'

Changes in **SharpeR** Version 1.2.1 (2020-02-06)

- CRAN fix for warnings about ellipsis.

Changes in **SharpeR** Version 1.2.0 (2018-10-07)

- move github figures to location CRAN understands
- be smarter about S3 classes: do not redefine summary and print.
- add bias and variance from Bao (2009).
- support estimation of higher order moments in `as.sr`, and expands methods for `se` and confidence interval computations.
- incorporate higher order methods into one sample `sr` tests.

Changes in **SharpeR** Version 1.1.0 (2016-03-14)

- fix `sr_vcov` on array input.
- add SRIC method.
- add SRIC to `print.sropt`.
- change `predint` output to matrix.

Changes in **SharpeR** Version 1.0.0 (2015-06-18)

- sane version numbers.
- unpaired k sample test of Sharpe.
- rely on same for unpaired 2 sample test.
- prediction intervals for Sharpe based on epsilon.
- more tests.

Changes in **SharpeR** Version 0.1501 (2014-12-06)

- fix inference of mark frequency from e.g. `xts` objects.
- add `rlambda`.

Changes in SharpeR Version 0.1401 (2014-01-05)

- fix second moment asymptotic covariance.
- add confidence distribution functions for sr, sr.opt.

Changes in SharpeR Version 0.1310 (2013-10-30)

- inverse second moment asymptotic covariance.

Changes in SharpeR Version 0.1309 (2013-09-20)

- spanning/hedging tests.
- sr equality test via callback variance covariance computation.
- split vignette in two.

Changes in SharpeR Version 0.1307 (2013-05-30)

- proper d.f. in sr objects with different nan fill.
- restore vignette.

SharpeR Initial Version 0.1306 (2013-05-21)

- put on CRAN

 sm_vcov

Compute variance covariance of 'Unified' Second Moment

Description

Computes the variance covariance matrix of sample mean and second moment.

Usage

```
sm_vcov(X,vcov.func=vcov,fit.intercept=TRUE)
```

Arguments

<code>X</code>	an $n \times p$ matrix of observed returns.
<code>vcov.func</code>	a function which takes an object of class <code>lm</code> , and computes a variance-covariance matrix. If equal to the string "normal", we assume multivariate normal returns.
<code>fit.intercept</code>	a boolean controlling whether we add a column of ones to the data, or fit the raw uncentered second moment.

Details

Given p -vector x , the 'unified' sample is the $p(p+3)/2$ vector of x stacked on top of $\text{vech}(xx^\top)$. Given n contemporaneous observations of p -vectors, stacked as rows in the $n \times p$ matrix X , this function computes the mean and the variance-covariance matrix of the 'unified' sample.

One may use the default method for computing covariance, via the `vcov` function, or via a 'fancy' estimator, like `sandwich:vcovHAC`, `sandwich:vcovHC`, *etc.*

Value

a list containing the following components:

<code>mu</code>	a $q = p(p+3)/2$ vector of the mean, then the vech'd second moment of the sample data
<code>Ohat</code>	the $q \times q$ estimated variance covariance matrix. Only the informative part is returned: one may assume a row and column of zeros in the upper left.
<code>n</code>	the number of rows in X .
<code>p</code>	the number of assets.

Note

This function will be deprecated in future releases of this package. Users should migrate at that time to a similar function in the MarkowitzR package.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Pav, S. E. "Asymptotic Distribution of the Markowitz Portfolio." 2013 <http://arxiv.org/abs/1312.0557>

See Also

[ism_vcov](#), [sr_vcov](#)

Examples

```
X <- matrix(rnorm(1000*3),ncol=3)
Sigmas <- sm_vcov(X)
Sigmas.n <- sm_vcov(X,vcov.func="normal")
Sigmas.n <- sm_vcov(X,fit.intercept=FALSE)

# make it fat tailed:
X <- matrix(rt(1000*3,df=5),ncol=3)
Sigmas <- sm_vcov(X)

if (require(sandwich)) {
  Sigmas <- sm_vcov(X,vcov.func=vcovHC)
```

```

}

# add some autocorrelation to X
Xf <- filter(X,c(0.2),"recursive")
colnames(Xf) <- colnames(X)
Sigmas <- sm_vcov(Xf)

if (require(sandwich)) {
  Sigmas <- sm_vcov(Xf,vcov.func=vcovHAC)
}

```

sr *Create an 'sr' object.*

Description

Spawns an object of class sr.

Usage

```

sr(
  sr,
  df,
  c0 = 0,
  ope = 1,
  rescal = sqrt(1/(df + 1)),
  epoch = "yr",
  cumulants = NULL
)

```

Arguments

sr	a Sharpe ratio statistic.
df	the degrees of freedom of the equivalent t-statistic.
c0	the 'risk-free' or 'disastrous' rate of return. this is assumed to be given in the same units as x, <i>not</i> in 'annualized' terms.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
rescal	the rescaling parameter.
epoch	the string representation of the 'epoch', defaulting to 'yr'.

cumulants an optional array of the higher order cumulants of the returns distribution. The first element shall be the skew; the second the excess kurtosis. Up to the sixth cumulant can be given. Higher order approximations for the moments of the Sharpe ratio can be computed based on these cumulants.

Details

The `sr` class contains information about a rescaled t-statistic. The following are list attributes of the object:

sr The Sharpe ratio statistic.

df The d.f. of the equivalent t-statistic.

c0 The drag 'risk free rate' used.

ope The 'observations per epoch'.

rescal The rescaling parameter.

epoch The string name of the 'epoch'.

The stored Sharpe statistic, `sr` is equal to the t-statistic times $rescal * sqrt(ope)$.

For the most part, this constructor should *not* be called directly, rather `as.sr` should be called instead to compute the Sharpe ratio.

Value

a list cast to class `sr`.

Note

2FIX: allow rownames?

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

See Also

[reannualize](#)

[as.sr](#)

Other `sr`: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [summary.sr](#)

Examples

```
# roll your own.
ope <- 253
zeta <- 1.0
n <- 3 * ope
rvs <- rsr(1,n,zeta,ope=ope)
roll.own <- sr(sr=rvs,df=n-1,ope=ope,rescal=sqrt(1/n))
# put a bunch in. naming becomes a problem.
rvs <- rsr(5,n,zeta,ope=ope)
roll.own <- sr(sr=rvs,df=n-1,ope=ope,rescal=sqrt(1/n))
```

sric

*Sharpe Ratio Information Coefficient***Description**

Computes the Sharpe Ratio Information Coefficient of Paulsen and Soehl, an asymptotically unbiased estimate of the out-of-sample Sharpe of the in-sample Markowitz portfolio.

Usage

```
sric(z.s)
```

Arguments

`z.s` an object of type `sropt`

Details

Let X be an observed $T \times k$ matrix whose rows are i.i.d. normal. Let μ and Σ be the sample mean and sample covariance. The Markowitz portfolio is

$$w = \Sigma^{-1}\mu,$$

which has an in-sample Sharpe of $\zeta = \sqrt{\mu^\top \Sigma^{-1} \mu}$.

The *Sharpe Ratio Information Criterion* is defined as

$$SRIC = \zeta - \frac{k-1}{T\zeta}.$$

The expected value (over draws of X and of future returns) of the *SRIC* is equal to the expected value of the out-of-sample Sharpe of the (in-sample) portfolio w (again, over the same draws.)

Value

The Sharpe Ratio Information Coefficient.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Paulsen, D., and Soehl, J. "Noise Fit, Estimation Error, and Sharpe Information Criterion." arxiv preprint (2016): <http://arxiv.org/abs/1602.06186>

See Also

Other sropt Hotelling: [inference\(\)](#)

Examples

```
# generate some sropts
nfac <- 3
nyr <- 5
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("fix seed")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
srv <- sric(asro)
```

sropt

Create an 'sropt' object.

Description

Spawns an object of class sropt.

Usage

```
sropt(z.s, df1, df2, drag = 0, ope = 1, epoch = "yr", T2 = NULL)
```

Arguments

<code>z.s</code>	an optimum Sharpe ratio statistic.
<code>df1</code>	the number of assets in the portfolio.
<code>df2</code>	the number of observations.
<code>drag</code>	the 'drag' term, c_0/R . defaults to 0. It is assumed that drag has been annualized, <i>i.e.</i> has been multiplied by \sqrt{ope} . This is in contrast to the c_0 term given to sr .

ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
epoch	the string representation of the 'epoch', defaulting to 'yr'.
T2	the Hotelling T^2 statistic. If not given, it is computed from the given information.

Details

The `sropt` class contains information about a rescaled T^2 -statistic. The following are list attributes of the object:

sropt The (optimal) Sharpe ratio statistic.

df1 The number of assets.

df2 The number of observations.

drag The drag term, which is the 'risk free rate' divided by the maximum risk.

ope The 'observations per epoch'.

epoch The string name of the 'epoch'.

For the most part, this constructor should *not* be called directly, rather `as.sropt` should be called instead to compute the needed statistics.

Value

a list cast to class `sropt`, with the following attributes:

sropt the optimal Sharpe statistic.

df1 the number of assets.

df2 the number of observed vectors.

drag the input drag term.

ope the input ope term.

epoch the input epoch term.

T2 the Hotelling T^2 statistic.

Note

2FIX: allow rownames?

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also[reannualize](#)[as.sropt](#)

Other sropt: [as.sropt\(\)](#), [confint.sr\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt_test\(\)](#)

Examples

```
# roll your own.
ope <- 253
zeta.s <- 1.0
df1 <- 10
df2 <- 6 * ope
set.seed(as.integer(charToRaw("fix seed")))
rvs <- rsropt(1,df1,df2,zeta.s,ope,drag=0)
roll.own <- sropt(z.s=rvs,df1,df2,drag=0,ope=ope)
print(roll.own)
# put a bunch in. naming becomes a problem.
rvs <- rsropt(5,df1,df2,zeta.s,ope,drag=0)
roll.own <- sropt(z.s=rvs,df1,df2,drag=0,ope=ope)
print(roll.own)
```

sropt_test

*test for optimal Sharpe ratio***Description**

Performs one sample tests of Sharpe ratio of the Markowitz portfolio.

Usage

```
sropt_test(X,alternative=c("greater", "two.sided", "less"),
           zeta.s=0,ope=1,conf.level=0.95)
```

Arguments

<code>X</code>	a (non-empty) numeric matrix of data values, each row independent, each column representing an asset, or an object of class <code>sropt</code> .
<code>alternative</code>	a character string specifying the alternative hypothesis, must be one of "two.sided", "greater" (default) or "less". You can specify just the initial letter.
<code>zeta.s</code>	a number indicating the null hypothesis value.
<code>ope</code>	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of <code>ope</code> per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
<code>conf.level</code>	confidence level of the interval. (not used yet)

Details

Suppose x_i are n independent draws of a q -variate normal random variable with mean μ and covariance matrix Σ . This code tests the hypothesis

$$H_0 : \mu^\top \Sigma^{-1} \mu = \delta_0^2$$

The default alternative hypothesis is the one-sided

$$H_1 : \mu^\top \Sigma^{-1} \mu > \delta_0^2$$

but this can be set otherwise.

Note there is no 'drag' term here since this represents a linear offset of the population parameter.

Value

A list with class "htest" containing the following components:

statistic	the value of the T^2 -statistic.
parameter	a list of the degrees of freedom for the statistic.
p.value	the p-value for the test.
conf.int	a confidence interval appropriate to the specified alternative hypothesis. NYI.
estimate	the estimated optimal Sharpe, annualized
null.value	the specified hypothesized value of the optimal Sharpe.
alternative	a character string describing the alternative hypothesis.
method	a character string indicating what type of test was performed.
data.name	a character string giving the name(s) of the data.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

See Also

[reannualize](#)

[sr_test](#), [t.test](#).

Other sropt: [as.sropt\(\)](#), [confint.sr\(\)](#), [dsropt\(\)](#), [is.sropt\(\)](#), [pco_sropt\(\)](#), [power.sropt_test\(\)](#), [reannualize\(\)](#), [sropt](#)

Examples

```
# test for uniformity
pvs <- replicate(128, { x <- sropt_test(matrix(rnorm(1000*4), ncol=4), alternative="two.sided")
  x$p.value })
plot(ecdf(pvs))
abline(0, 1, col='red')
```

```
# input a spropt objects:
nfac <- 5
nyr <- 10
ope <- 253
# simulations with no covariance structure.
# under the null:
set.seed(as.integer(charToRaw("be determinstic")))
Returns <- matrix(rnorm(ope*nyr*nfac,mean=0,sd=0.0125),ncol=nfac)
asro <- as.sropt>Returns,drag=0,ope=ope)
stest <- spropt_test(asro,alternative="two.sided")
```

sr_bias

sr_bias.

Description

Computes the asymptotic bias of the sample Sharpe ratio based on moments.

Usage

```
sr_bias(snr, n, cumulants, type = c("simple", "second_order"))
```

Arguments

snr	the population Signal Noise ratio. Often one will use the population estimate instead.
n	the sample size that the Shapre ratio is observed on.
cumulants	a vector of the third through fourth, or the third through seventh population cumulants of the random variable. More terms are needed for the higher accuracy approximation.
type	determines the order of accuracy of the bias approximation. Takes values of simple We compute the simple approximation using only the skewness and excess kurtosis. second_order We compute the more accurate approximation, given by Bao, which is accurate to $o(n^{-2})$.

Details

The sample Sharpe ratio has bias of the form

$$B = \left(\frac{3}{4n} + 3\frac{\gamma_2}{8n} \right) \zeta - \frac{1}{2n} \gamma_1 + o(n^{-3/2}),$$

where ζ is the population Signal Noise ratio, n is the sample size, γ_1 is the population skewness, and γ_2 is the population excess kurtosis. This form of the bias appears as Equation (5) in Bao, which claims an accuracy of only $o(n^{-1})$. The author believes this approximation is slightly more accurate.

A more accurate form is given by Bao (Equation (3)) as

$$B = \frac{3\zeta}{4n} + \frac{49\zeta}{32n^2} - \gamma_1 \left(\frac{1}{2n} + \frac{3}{8n^2} \right) + \gamma_2 \zeta \left(\frac{3}{8n} - \frac{15}{32n^2} \right) + \frac{3\gamma_3}{8n^2} - \frac{5\gamma_4\zeta}{16n^2} - \frac{5\gamma_1^2\zeta}{4n^2} + \frac{105\gamma_2^2\zeta}{128n^2} - \frac{15\gamma_1\gamma_2}{16n^2} + o(n^{-2}),$$

where γ_3 through γ_5 are the fifth through seventh cumulants of the error term.

Value

the approximate bias of the Sharpe ratio. The bias is the expected value of the sample Sharpe minus the Signal Noise ratio.

Note

much of the code is adapted from Gauss code provided by Yong Bao.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Bao, Yong. "Estimation Risk-Adjusted Sharpe Ratio and Fund Performance Ranking Under a General Return Distribution." *Journal of Financial Econometrics* 7, no. 2 (2009): 152-173. <https://doi.org/10.1093/jjfinec/nbn022>

See Also

[sr_variance](#)

Examples

```
# bias under normality:
sr_bias(1, 100, rep(0,2), type='simple')
sr_bias(1, 100, rep(0,5), type='second_order')

# plugging in sample estimates
x <- rnorm(1000)
n <- length(x)
mu <- mean(x)
sdv <- sd(x)
snr <- mu / sdv
# these are not great estimates, but close enough:
sku <- mean((x-mu)^3) / sdv^3
kur <- (mean((x-mu)^4) / sdv^4) - 4
sr_bias(snr, n, c(sku,kur), type='simple')
```

sr_equality_test Paired test for equality of Sharpe ratio

Description

Performs a hypothesis test of equality of Sharpe ratios of p assets given paired observations.

Usage

```
sr_equality_test(X,type=c("chisq","F","t"),
                 alternative=c("two.sided","less","greater"),
                 contrasts=NULL,
                 vcov.func=vcov)
```

Arguments

X	an $n \times p$ matrix of paired observations.
type	which approximation to use. "chisq" is preferred when the returns are non-normal, but the approximation is asymptotic. the "t" test is only supported when $k = 1$.
alternative	a character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". You can specify just the initial letter. This is only relevant for the "t" test. "greater" corresponds to $H_a : E s > 0$.
contrasts	an $k \times p$ matrix of the contrasts
vcov.func	a function which takes a model of class lm (one of the form $x \sim 1$), and produces a variance-covariance matrix. The default is <code>vcov</code> , which produces a 'vanilla' estimate of covariance. Other sensible options are <code>vcovHAC</code> from the <code>sandwich</code> package.

Details

Given n *i.i.d.* observations of the excess returns of p strategies, we test

$$H_0 : \frac{\mu_i}{\sigma_i} = \frac{\mu_j}{\sigma_j}, 1 \leq i < j \leq p$$

using the method of Wright, et. al.

More generally, a matrix of constrasts, E can be given, and we can test

$$H_0 : E s = 0,$$

where s is the vector of Sharpe ratios of the p strategies.

When E consists of a single row (a single contrast), as is the case when the default contrasts are used and only two strategies are compared, then an approximate t-test can be performed against the alternative hypothesis $H_a : E s > 0$

Both chi-squared and F- approximations are supported; the former is described by Wright. *et. al.*, the latter by Leung and Wong.

Value

Object of class `htest`, a list of the test statistic, the size of X , and the method noted.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

- Sharpe, William F. "Mutual fund performance." *Journal of business* (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>
- Wright, J. A., Yam, S. C. P., and Yung, S. P. "A note on the test for the equality of multiple Sharpe ratios and its application on the evaluation of iShares." *J. Risk*. to appear. <http://www.risk.net/journal-of-risk/technical-paper/2340067/a-test-for-the-equality-of-multiple-sharpe-ratios>
- Leung, P.-L., and Wong, W.-K. "On testing the equality of multiple Sharpe ratios, with application on the evaluation of iShares." *J. Risk* 10, no. 3 (2008): 15-30. http://papers.ssrn.com/sol3/papers.cfm?abstract_id=907270
- Memmel, C. "Performance hypothesis testing with the Sharpe ratio." *Finance Letters* 1 (2003): 21-23.
- Ledoit, O., and Wolf, M. "Robust performance hypothesis testing with the Sharpe ratio." *Journal of Empirical Finance* 15, no. 5 (2008): 850-859. http://www.ledoit.net/jef2008_abstract.htm
- Lo, Andrew W. "The statistics of Sharpe ratios." *Financial Analysts Journal* 58, no. 4 (2002): 36-52. <http://ssrn.com/paper=377260>

See Also

[sr_test](#)

Other `sr`: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr](#), [summary.sr](#)

Examples

```
# under the null
set.seed(1234)
rv <- sr_equality_test(matrix(rnorm(500*5),ncol=5))

# under the alternative (but with identity covariance)
ope <- 253
nyr <- 10
nco <- 5
set.seed(909)
rets <- 0.01 * sapply(seq(0,1.7/sqrt(ope),length.out=nco),
  function(mu) { rnorm(ope*nyr,mean=mu,sd=1) })
rv <- sr_equality_test(rets)

# using real data
if (require(xts)) {
  data(stock_returns)
```

```

pvs <- sr_equality_test(stock_returns)
}

# test for uniformity
pvs <- replicate(1024, { x <- sr_equality_test(matrix(rnorm(400*5), 400, 5), type="chisq")
                                     x$p.value })

plot(ecdf(pvs))
abline(0, 1, col='red')

if (require(sandwich)) {
  set.seed(as.integer(charToRaw("0b2fd4e9-3bdf-4e3e-9c75-25c6d18c331f")))
  n.manifest <- 10
  n.latent <- 4
  n.day <- 1024
  snr <- 0.95
  la_A <- matrix(rnorm(n.day*n.latent), ncol=n.latent)
  la_B <- matrix(runif(n.latent*n.manifest), ncol=n.manifest)
  latent.rets <- la_A %*% la_B
  noise.rets <- matrix(rnorm(n.day*n.manifest), ncol=n.manifest)
  some.rets <- snr * latent.rets + sqrt(1-snr^2) * noise.rets
  # naive vcov
  pvs0 <- sr_equality_test(some.rets)
  # HAC vcov
  pvs1 <- sr_equality_test(some.rets, vcov.func=vcovHAC)
  # more elaborately:
  pvs <- sr_equality_test(some.rets, vcov.func=function(amod) {
vcovHAC(amod, prewhite=TRUE) })
}

```

sr_test

test for Sharpe ratio

Description

Performs one and two sample tests of Sharpe ratio on vectors of data.

Usage

```

sr_test(
  x,
  y = NULL,
  alternative = c("two.sided", "less", "greater"),
  zeta = 0,
  ope = 1,
  paired = FALSE,
  conf.level = 0.95,

```



```

type = c("exact", "t", "Z", "Mertens", "Bao"),
...
)

```

Arguments

x	a (non-empty) numeric vector of data values, or an object of class <code>sr</code> , containing a scalar sample Sharpe estimate.
y	an optional (non-empty) numeric vector of data values, or an object of class <code>sr</code> , containing a scalar sample Sharpe estimate. Only an unpaired test can be performed when at least one of x and y are of class <code>sr</code>
alternative	a character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". You can specify just the initial letter.
zeta	a number indicating the null hypothesis offset value, the S value.
ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.
paired	a logical indicating whether you want a paired test.
conf.level	confidence level of the interval.
type	which method to apply.
...	further arguments to be passed to or from methods.

Details

Given n observations x_i from a normal random variable, with mean μ and standard deviation σ , tests

$$H_0 : \frac{\mu}{\sigma} = S$$

against two or one sided alternatives.

Can also perform two sample tests of Sharpe ratio. For paired observations x_i and y_i , tests

$$H_0 : \frac{\mu_x}{\sigma_x} = \frac{\mu_u}{\sigma_y}$$

against two or one sided alternative, via [sr_equality_test](#).

For unpaired (and independent) observations, tests

$$H_0 : \frac{\mu_x}{\sigma_x} - \frac{\mu_u}{\sigma_y} = S$$

against two or one-sided alternatives via the ϵ distribution.

The one sample test admits a number of different methods:

exact The default, which is only exact when returns are normal, based on inverting the non-central t distribution.

- t** Uses the Johnson Welch approximation to the standard error, centered around the sample value.
- Z** Uses the Johnson Welch approximation to the standard error, performing a simple correction for the bias of the Sharpe ratio based on Miller and Gehr formula.
- Mertens** Uses the Mertens higher order approximation to the standard error, centered around the sample value.
- Bao** Uses the Bao higher order approximation to the standard error, performing a higher order correction for the bias of the Sharpe ratio.

See [confint.sr](#) for more information on these types

Value

A list with class "htest" containing the following components:

statistic	the value of the t- or Z-statistic.
parameter	the degrees of freedom for the statistic.
p.value	the p-value for the test.
conf.int	a confidence interval appropriate to the specified alternative hypothesis. NYI for some cases.
estimate	the estimated Sharpe or difference in Sharpes depending on whether it was a one-sample test or a two-sample test. Annualized
null.value	the specified hypothesized value of the Sharpe or difference of Sharpes depending on whether it was a one-sample test or a two-sample test.
alternative	a character string describing the alternative hypothesis.
method	a character string indicating what type of test was performed.
data.name	a character string giving the name(s) of the data.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

Pav, Steven. "Inference on the Sharpe ratio via the upsilon distribution." Arxiv (2015). <http://arxiv.org/abs/1505.00829>

See Also

[reannualize](#)

[sr_equality_test](#), [sr_unpaired_test](#), [t.test](#).

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr](#), [summary.sr](#)

Examples

```

# should reject null
x <- sr_test(rnorm(1000,mean=0.5,sd=0.1),zeta=2,ope=1,alternative="greater")
x <- sr_test(rnorm(1000,mean=0.5,sd=0.1),zeta=2,ope=1,alternative="two.sided")
# should not reject null
x <- sr_test(rnorm(1000,mean=0.5,sd=0.1),zeta=2,ope=1,alternative="less")

# test for uniformity
pvs <- replicate(128,{ x <- sr_test(rnorm(1000),ope=253,alternative="two.sided")
                       x$p.value })

plot(ecdf(pvs))
abline(0,1,col='red')
# testing an object of class sr
asr <- as.sr(rnorm(1000,1 / sqrt(253)),ope=253)
checkit <- sr_test(asr,zeta=0)

```

sr_unpaired_test	<i>test for equation on unpaired Sharpe ratios</i>
------------------	--

Description

Performs hypothesis tests on a single equation on k independent samples of Sharpe ratio.

Usage

```

sr_unpaired_test(
  srs,
  contrasts = NULL,
  null.value = 0,
  alternative = c("two.sided", "less", "greater"),
  ope = NULL,
  conf.level = 0.95
)

```

Arguments

srs	a (non-empty) list of objects of class sr, each containing a scalar sample Sharpe estimate. Or a single object of class sr with multiple Sharpe estimates. If the sr objects have different annualizations (ope parameters), a warning is thrown, since it is presumed that the contrasts all have the same units, but the test proceeds.
contrasts	an array of the contrasts, the a_j values. Defaults to $c(1, -1, 1, \dots)$.
null.value	the constant null value, the b . Defaults to 0.
alternative	a character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". You can specify just the initial letter.

ope	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of ope per epoch. The default value is to take the same ope from the input srs object, if it is unambiguous. Otherwise, it defaults to 1, with a warning thrown.
conf.level	confidence level of the interval.

Details

For $1 \leq j \leq k$, suppose you have n_j observations of a normal random variable with mean μ_j and standard deviation σ_j , with all observations independent. Given constants a_j and value b , this code tests the null hypothesis

$$H_0 : \sum_j a_j \frac{\mu_j}{\sigma_j} = b$$

against two or one sided alternatives.

Value

A list with class "htest" containing the following components:

statistic	NULL here.
parameter	a list of epsilon parameters.
p.value	the p-value for the test.
conf.int	a confidence interval appropriate to the specified alternative hypothesis.
estimate	the estimated equation value, just the weighted sum of the sample Sharpe ratios. Annualized
null.value	the specified hypothesized value of the sum of Sharpes.
alternative	a character string describing the alternative hypothesis.
method	a character string indicating what type of test was performed.
data.name	a character string giving the name(s) of the data.

Note

This code is based on the 'epsilon' code from `sadists`, which may be inaccurate for a large number of series. Take caution when applying this test. File a bug report if you are negatively impacted.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

- Sharpe, William F. "Mutual fund performance." *Journal of business* (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>
- Pav, Steven. "Inference on the Sharpe ratio via the epsilon distribution." *Arxiv* (2015). <http://arxiv.org/abs/1505.00829>

See Also

[sr_equality_test](#), [sr_test](#), [t.test](#).

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_vcov\(\)](#), [sr](#), [summary.sr](#)

Examples

```
# basic usage
set.seed(as.integer(charToRaw("set the seed")))
# default contrast is 1,-1,1,-1,1,-1
etc <- sr_unpaired_test(as.sr(matrix(rnorm(1000*6,mean=0.02,sd=0.1),ncol=6)))
print(etc)

etc <- sr_unpaired_test(as.sr(matrix(rnorm(1000*4,mean=0.0005,sd=0.01),ncol=4)),
  alternative='greater')
print(etc)

etc <- sr_unpaired_test(as.sr(matrix(rnorm(1000*4,mean=0.0005,sd=0.01),ncol=4)),
  contrasts=c(1,1,1,1),null.value=-0.1,alternative='greater')
print(etc)

inp <- list(as.sr(rnorm(500)),as.sr(runif(200)-0.5),
  as.sr(rnorm(30)),as.sr(rnorm(100)))
etc <- sr_unpaired_test(inp)

inp <- list(as.sr(rnorm(500)),as.sr(rnorm(100,mean=0.2,sd=1)))
etc <- sr_unpaired_test(inp,contrasts=c(1,1),null.value=0.2)
etc$conf.int
```

 sr_variance

 sr_variance .

Description

Computes the variance of the sample Sharpe ratio.

Usage

```
sr_variance(snr, n, cumulants)
```

Arguments

snr	the population Signal Noise ratio. Often one will use the population estimate instead.
n	the sample size that the Sharpe ratio is observed on.
cumulants	a vector of the third through fourth, or the third through seventh population cumulants of the random variable. More terms are needed for the higher accuracy approximation.

Details

The sample Sharpe ratio has variance of the form

$$V = \frac{1}{n} \left(1 + \frac{\zeta^2}{2} \right) + \frac{1}{n^2} \left(\frac{19\zeta^2}{8} + 2 \right) - \gamma_1 \zeta \left(\frac{1}{n} + \frac{5}{2n^2} \right) + \gamma_2 \zeta^2 \left(\frac{1}{4n} + \frac{3}{8n^2} \right) + \frac{5\gamma_3 \zeta}{4n^2} + \gamma_1^2 \left(\frac{7}{4n^2} - \frac{3\zeta^2}{2n^2} \right) + \frac{39\gamma_2^2 \zeta^2}{32n^2} - \frac{15\gamma_1}{4n}$$

where ζ is the population Signal Noise ratio, n is the sample size, γ_1 is the population skewness, and γ_2 is the population excess kurtosis, and γ_3 through γ_5 are the fifth through seventh cumulants of the error term. This form of the variance appears as Equation (4) in Bao.

Value

the variance of the sample statistic.

Author(s)

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References

Bao, Yong. "Estimation Risk-Adjusted Sharpe Ratio and Fund Performance Ranking Under a General Return Distribution." *Journal of Financial Econometrics* 7, no. 2 (2009): 152-173. <https://doi.org/10.1093/jfinec/nbn022>

See Also

[sr_bias](#).

Examples

```
# variance under normality:
sr_variance(1, 100, rep(0,5))
```

sr_vcov

Compute variance covariance of Sharpe Ratios.

Description

Computes the variance covariance matrix of sample Sharpe ratios.

Usage

```
sr_vcov(X, vcov.func=vcov, ope=1)
```

Arguments

<code>X</code>	an $n \times p$ matrix of observed returns. If not a matrix, but a numeric of length n , then it is coerced into a $n \times 1$ matrix.
<code>vcov.func</code>	a function which takes an object of class <code>lm</code> , and computes a variance-covariance matrix.
<code>ope</code>	the number of observations per 'epoch'. For convenience of interpretation, The Sharpe ratio is typically quoted in 'annualized' units for some epoch, that is, 'per square root epoch', though returns are observed at a frequency of <code>ope</code> per epoch. The default value is 1, meaning the code will not attempt to guess what the observation frequency is, and no annualization adjustments will be made.

Details

Given n contemporaneous observations of p returns streams, this function estimates the asymptotic variance covariance matrix of the vector of sample Sharpes, $[\zeta_1, \zeta_2, \dots, \zeta_p]$

One may use the default method for computing covariance, via the `vcov` function, or via a 'fancy' estimator, like `sandwich:vcovHAC`, `sandwich:vcovHC`, *etc.*

This code first estimates the covariance of the $2p$ vector of the vector x stacked on its Hadamard square, x^2 . This is then translated back to a variance covariance on the vector of sample Sharpe ratios via the Delta method.

Value

a list containing the following components:

<code>SR</code>	a vector of (annualized) Sharpe ratios.
<code>Ohat</code>	a $p \times p$ variance covariance matrix.
<code>p</code>	the number of assets.

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

- Sharpe, William F. "Mutual fund performance." *Journal of business* (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>
- Lo, Andrew W. "The statistics of Sharpe ratios." *Financial Analysts Journal* 58, no. 4 (2002): 36-52. <http://ssrn.com/paper=377260>

See Also

[reannualize](#)

sr-distribution functions, [dsr](#)

Other sr: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambda\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr](#), [summary.sr](#)

Examples

```
X <- matrix(rnorm(1000*3),ncol=3)
colnames(X) <- c("ABC","XYZ","WORM")
Sigmas <- sr_vcov(X)
# make it fat tailed:
X <- matrix(rt(1000*3,df=5),ncol=3)
Sigmas <- sr_vcov(X)

if (require(sandwich)) {
  Sigmas <- sr_vcov(X,vcov.func=vcovHC)
}

# add some autocorrelation to X
Xf <- filter(X,c(0.2),"recursive")
colnames(Xf) <- colnames(X)
Sigmas <- sr_vcov(Xf)

if (require(sandwich)) {
  Sigmas <- sr_vcov(Xf,vcov.func=vcovHAC)
}

# should run for a vector as well
X <- rnorm(1000)
SS <- sr_vcov(X)
```

stock_returns

Stock Returns Data

Description

Nineteen years of daily log returns on three stocks and an ETF.

Usage

```
data(stock_returns)
```

Format

An xts object with 4777 observations and 4 columns.

The columns are the daily log returns for the tickers IBM, AAPL, SPY and XOM, as sourced from Yahoo finance using the quantmod package. Daily returns span from January, 2000 through December, 2018. Returns are 'log returns', which are the differences of the logs of daily adjusted closing price series, as defined by Yahoo finance (thus presumably including adjustments for splits and dividends). Dates of observations are the date of the second close defining the return, not the first.

Note

The author makes no guarantees regarding correctness of this data.

Author(s)

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Source

Data were collected on October 2, 2019, from Yahoo finance using the quantmod package.

Examples

```
if (require(xts)) {
  data(stock_returns)
  as.sr(stock_returns)
}
```

summary.sr

Summarize a Sharpe, or (delta) optimal Sharpe object.

Description

Computes a 'summary' of an object, adding in some statistics.

Usage

```
## S3 method for class 'sr'
summary(object, ...)

## S3 method for class 'sropt'
summary(object, ...)
```

Arguments

object an object of class sr, sropt or del_sropt.
... additional arguments affecting the summary produced, though ignored here.

Details

Enhances an object of class sr, sropt or del_sropt to also include t- or T-statistics, p-values, and so on.

Value

When an `sr` object is input, the object cast to class `summary.sr` with some additional fields:

tval the equivalent t-statistic.

pval the p-value under the null.

serr the standard error of the Sharpe ratio.

When an `sropt` object is input, the object cast to class `summary.sropt` with some additional fields:

pval the p-value under the null.

SRIC the SRIC value, see [sric](#).

Author(s)

Steven E. Pav <shabbychef@gmail.com>

References

Sharpe, William F. "Mutual fund performance." Journal of business (1966): 119-138. <http://ideas.repec.org/a/ucp/jnlbus/v39y1965p119.html>

See Also

[print.sr](#).

Other `sr`: [as.sr\(\)](#), [confint.sr\(\)](#), [dsr\(\)](#), [is.sr\(\)](#), [plambdap\(\)](#), [power.sr_test\(\)](#), [predint\(\)](#), [print.sr\(\)](#), [reannualize\(\)](#), [se\(\)](#), [sr_equality_test\(\)](#), [sr_test\(\)](#), [sr_unpaired_test\(\)](#), [sr_vcov\(\)](#), [sr](#)

Examples

```
# Sharpe's 'model': just given a bunch of returns.
set.seed(1234)
asr <- as.sr(rnorm(253*3), ope=253)
summary(asr)
```

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